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## Microprogramming For Probability Distribution Sampling

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Microprogramming of special instructions for sampling of random variates from any probability distribution is a means of increasing sampling speed. The diversity of sampling techniques is narrowed to one general algorithm; conditional bit sampling. Conditional bit sampling, uses a high speed uniform random number generator based on feedback shift registers to sample one bit at a time. The probability of a bit being a one in the  $j$ -th position of a binary expanded variate is stored in a table of conditional probabilities. A comparison with the pseudorandom number yields a one or zero. The table of conditional probabilities is generated once and passed through an instruction to the microprogram which performs the sampling. One user instruction is issued for each variate returned.

### INTRODUCTION

Microprogramming, under the control of users, allows special purpose computers to be implemented through software changes (minor hardware changes to ROM) [1]. The user may design instructions particular to his application without great speed degradation [10]. A computer manufacturer may design and build machines with great flexibility and yet retain low cost.

Advances in control memory have made dynamic alteration of control memory possible [3]. The instruction repertoire may be changed under program control in order to adapt to a special purpose application.

Thus, microprogramming is emerging as a design tool within reach of the applications implementor.

Non-deterministic simulation requires sampling from various probability distributions. Probabilistic simulation with a digital computer is based on pseudorandom number generators [2, 13, 4]. Most non-uniform probability distributions are sampled using uniformly random numbers [11, 12]. Software routines and special simulation languages exist for a wide range of simulation models, however, sampling speed is moderate. To increase efficiency of sampling a user may implement special hardwired computers or microprogram a general purpose computer. Appropriate techniques for microprogrammed sampling from probability distributions presented here are based on shift register theory and conditional probability tables.

### UNIFORM RANDOM NUMBERS

A microprogram for generating a uniform random integer is given in figure 1 [4, 5]. Programmed on the Interdata 4, this instruction generates a uniformly random integer each time it is issued by a user [6, 7]. A memory table of length  $n+2$  is organized in the 16-bit user memory as shown in Figure 2 and passed via address to the microprogram.

The GFSR (generalized feedback shift register) algorithm is:

1.  $I \leftarrow I + 1$
2. IF  $I > n$  THEN  $I \leftarrow 1$

MODEL 4 MICRO CODE

```

                                ORG X'428'
0428 5004      RND   L   RAH,H(RND)
0429 3080      C     SB
042A 4453      L     MR4,MAR      SAVE TABL IN MR4.
042B 43A3      L     MR3,MDR      SAVE I IN MR3.
042C 5802      L     AR,X'2'
042D C550      A     MAR,MAR,NF+NC
042E 3100      C     MR
042F C050      A     MR0,MAR,NF+NC  GET S,N: LOC TABL+2.
                                SAVE TABL+4 IN MR0.
0430 58FF      L     AR,X'FF'
0431 82A3      N     MR2,MDR,NF    MASK-OFF S; N IN MR2.
0432 48AF      L     AR,MDR,CS
0433 91FF      N     MR1,X'FF'     CROSS SHIFT AND MASK-C
0434 4833      L     AR,MR3        COMPARE I:N.
0435 E824      S     AR,MR2,NC
0436 1138      B     L,OK          BRANCH IF I<N.
0437 5300      L     MR3,X'00'     OTHERWISE, SET I=0
0438 4813      OK    L     AR,MR1
0439 C130      A     MR1,MR3,NF+NC  J=I+S
043A 4813      L     AR,MR1
043B E224      S     MR2,MR2,NC    (MR2)=J-N.
043C 113E      B     L,OK2        BRANCH IF J<N.
043D 4123      L     MR1,MR2     OTHERWISE, SET J=J-N.
043E 4818      OK2  L     AR,MR1,SL+NC  D1=2*J.
043F C500      A     MAR,MRO,NF+NC  LOC TABL+D1+4.
0440 3100      C     MR          GET TABL(J).
0441 4838      L     AR,MR3,SL+NC  D2=2*I.
0442 C500      A     MAR,MRO,NF+NC  LOC TABL+D2+4.
0443 48A3      L     AR,MDR      SAVE TABL(J).
0444 3100      C     MR          GET TABL(I).
0445 AEA3      X     YD,MDR,NF     EX-OR TABL(J)+TABL(I).
0446 4AE3      L     MDR,YD
0447 3200      C     MW          RESULT IN TABL(I).
0448 5801      L     AR,X'01'     INCREMENT.
0449 CA30      A     MDR,MR3,NF+NC  I=I+1.
044A 4543      L     MAR,MR4
044B 3200      C     MW          NEW I IN LOC TABL.
044C 4563      L     MAR,LOC
044D 066B      D     LOC,LOC,P2N
                                END

```

Fig. 1--Interdata 4 microprogram for GFSR. Execute, RNG REG1, TABLE. The resulting random integer is returned to register REG1, and the address of TABLE locates the memory table.

3.  $j \leftarrow j + S$
4. IF  $j > n$  THEN  $j \leftarrow j - n$
5.  $W_I \leftarrow W_I \oplus W_j$ .

GFSR must be initialized by setting  $I = 0$ , and generating an initial set of words,  $W_I$ , figure 2. Note that  $I$  runs from 1 to  $n$  in the algorithm, but to facilitate addressing runs from 0 to  $n-1$  in the microprogram.

The GFSR algorithm is fast and produces arbitrarily long period sequences [4]. It is based on composite feedback shift

register sequences which in turn are based on primitive trinomials over a binary field:  $X^n + X^s + 1$  [4, 13]. For example,  $X^4 + X^1 + 1$ , yields a basic binary sequence which may be combined with other delayed basic sequences to produce a sequence of random integers, figure 3. The recurrence relation used to produce a basic sequence is  $a_I = a_{I-n+s} \oplus a_{I-n}$ , or  $a_I = a_{I-3} \oplus a_{I-4}$ , in figure 3. The parameter,  $n$  and  $s$  are

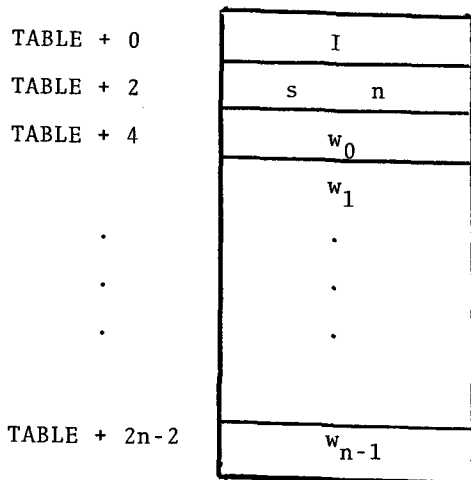


Fig. 2--Organization of TABLE for micro-program RND. GFSR uses  $x^n + x^s + 1$  by exclusive-or of  $w_n$  with  $w_{n+s}$ .

selected from a list of primitive trinomials and by considering memory and period requirements [4].

Statistical properties are guaranteed when relative delay between basic sequences is great [4]. An initialization routine may be run once for a particular trinomial,  $X^n + X^s + 1$ , and stored in a spare table in memory. The rotating table of figure 2 is altered during use and must be reset (by copying from the spare table) if the sequence of random numbers is to be reproduced.

One instruction (D4<sub>16</sub> in Interdata 4) is convenient for the user and when micro-programmed becomes very fast (few memory cycles). Production of several hundred thousand numbers per second is possible with firmware.

#### SAMPLING FROM DISTRIBUTIONS

Distribution sampling is an art [12]. Special subroutines based on analytic inversion, rational approximation, rejection, and composition have been implemented in software. For example, to sample from  $p(x) = 3X^2$ ,  $0 \leq X \leq 1$ , merely solve for X in the inversion formula;  $R = X^3$ . Here, R is a uniform random number,  $0 \leq R \leq 1$  and X is the positive cube root of R.

Microprogramming of each special case distribution would involve large amounts of control storage and microprogramming effort. Instead, a general method applicable to all continuous distributions is implemented as a microprogram. The resulting microcoded instruction operates on a table of conditional probabilities and returns one sampled integer each time it is issued.

The conditional probability table is created by computing a ratio of areas as shown in figure 4, [8]. Briefly, the binary number to be sampled is constructed by sampling for one bit at a time to get  $SX_1 X_2 \cdot X_3 X_4 \dots$ . The sign bit, S, and each bit,  $X_i$   $i=1,2,\dots$  is sampled by comparing a random number (produced by GFSR) with the conditional probability. The conditional probability table for  $p(x)$  is generated by the process given for variates with binary point between  $x_2$  and  $x_3$ , without loss of generality.

$$\begin{aligned} \Phi(Z) &= \int_0^Z p(x) dx \text{ [assume S sampled separately]} \\ P(x_1=1) &= \frac{2[100.00\dots] - \Phi(10.00\dots)}{\Phi(10.00\dots)} \\ P(x_1=0) &= 1 - P(x_1 = 1) \\ P(x_2=1|x_1 = 0) &= \frac{\Phi(01.111\dots=10.00\dots) - \Phi(01.00\dots)}{\Phi(01.11\dots=10.00\dots) - \Phi(00.00\dots)} \\ &= \frac{2[\Phi(2) - \Phi(1)]}{2[\Phi(2) - \Phi(0)]} \end{aligned}$$

Similarly,

$$\begin{aligned} P(x_2 = 1|x_1 = 1) &= \frac{P(x_1=1 \text{ and } x_2=1)}{P(x_1 = 1)} \\ &= \frac{\Phi(4) - \Phi(3)}{\Phi(4) - \Phi(2)} \end{aligned}$$

In general, for the magnitude portion,  $\bar{X}_j = X_1 X_2 \cdot X_3 \dots X_j$ , it is possible to inductively obtain;

$$P(x_{j+1} = 1|\bar{x}_j) =$$



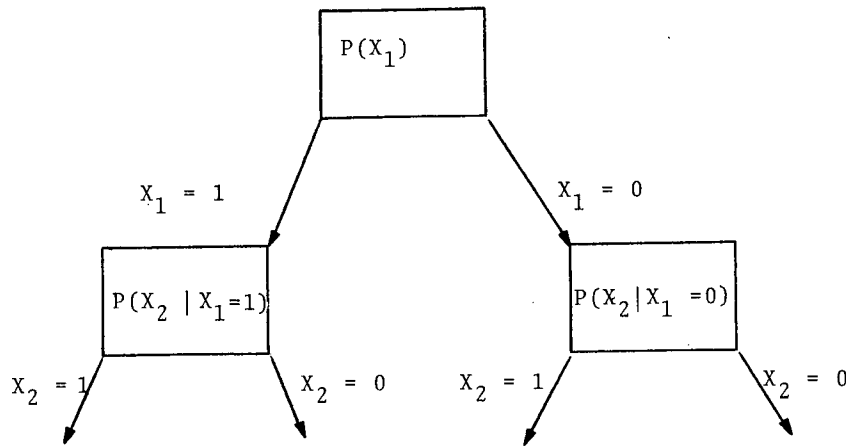


Fig. 5--A conditional probability tree for bit sampling.

$$\frac{\Phi(\bar{x}_j + 2^{2-j}) - \Phi(\bar{x}_j + 2^{1-j})}{\Phi(\bar{x}_j + 2^{2-j}) - \Phi(\bar{x}_j)} ;$$

$j = 1, 2, \dots$

$$P(x_1 = 1 | \bar{x}_0) = P(x_1 = 1)$$

Thus, a conditional probability tree as shown in figure 5 is created for bit sampling.

Substitution of  $Z_j = \bar{x}_j + 2^{2-j}$  into the expression for  $P(x_{j+1} = 1 | \bar{x}_j)$  yields:

$$P(x_{j+1} = 1 | \bar{x}_j) = \frac{1}{2} \frac{[\Phi(Z_j) - \Phi(Z_j - 2^{1-j})]/2^{1-j}}{[\Phi(Z_j) - \Phi(Z_j - 2^{2-j})]/2^{2-j}}$$

and,

$$\lim_{j \rightarrow \infty} P(x_{j+1} = 1 | \bar{x}_j) = \frac{1}{2} \frac{d\Phi/dZ}{d\Phi/dZ} = \frac{1}{2}$$

$j \rightarrow \infty$

Thus, if  $\Phi$  is differentiable, then the occurrence of the  $j$ -th bit of  $X$  is uniformly distributed in the limit. This implies that a uniform random number may be appended to the low order bits of a random variate sampled from a differentiable distribution. The details of this technique can be found in [8], but suffice it to say that greater accuracy is obtained simply by augmenting the random variate

with a random number. The number of bits to be sampled by conditional bit sampling may be determined by observing the convergence of  $P(x_{j+1} | \bar{x}_j)$  to one-half. For example, for the normal distribution ( $\mu=0$ ,  $\sigma = 1$ ) convergence is very close to one-half after seven bits are sampled.

The binary point in  $SX_1 X_2 \dots X_3 \dots$  may be moved so that scaling to an integer is possible during sampling and then replaced after obtaining a sample. Also, increased accuracy is obtained by arbitrarily selecting the number of leading bits (to the left of point). Finally,  $S$  is sampled by weighing according to the fraction of samples which are negative/positive.

The total number of random numbers required is moderately high compared to specialized routines, but this is offset by the speed of GFSR. The CBS (conditional bit sampling) algorithm is:

1.  $I \leftarrow 0$
2.  $J \leftarrow 0$
3.  $M \leftarrow I + 2 \uparrow J$
4.  $I \leftarrow 2 * I$
5. IF RAND  $\leq$  TABLE(M) THEN  $I \leftarrow I + 1$ .
6.  $J \leftarrow J + 1$
7. IF  $J < L$  GOTO 3.
8. STOP I IS SAMPLED VARIATE

To normalize  $I$  in CBS to floating point number with  $K$  bits to the left of the binary point:

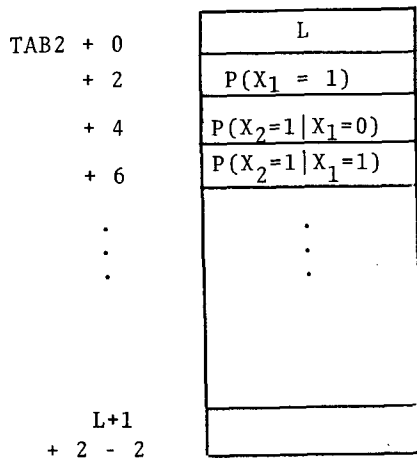


Fig. 6--Organization of CBS algorithm table, TAB2, containing conditional probabilities.

$$\text{Float } (i) + I/(2+(L-K))$$

The selection of L is determined by convergence of probability table, TABLE, which is linear array containing the tree of probabilities.

The CBS algorithm is easily microprogrammed as instruction, VAR REG1, TABLE, just as before in GFSR. A user-memory table, TABLE, is passed to the VAR microprogram. In microprogrammed machines which do not allow sub-microprograms to be called (transfer within control memory) a more complex instruction, VAR, REG1, TAB1, TAB2 is required:

REG1 = returned random variate

TAB1 = address of GFSR table

TAB2 = address of CBS table

Therefore, the RND micro-routine is coded as an in-line routine. The organization of TAB2 is given in figure 6.

#### DISCRETE DISTRIBUTION SAMPLING

Sampling from empirical or discrete variable distribution is more simplified than from the continuous. The limiting uniform distribution no longer applies, however, (not needed) and some of the conditional probabilities may even be zero.

To compute conditional probabilities, say for the Gamma distribution,

$$p(x) = \frac{x^{p-1} e^{-x/b}}{\Gamma(p) b^p}$$

$$p \geq 1, b > 0, 0 < X < \infty$$

an integration is performed over sections of p(x). To create a conditional probability table for, say the Poisson distribution,

$$p_\lambda(x) = \frac{e^{-\lambda} \lambda^x}{x!}$$

$$\lambda > 0, x = 0, 1, 2, \dots$$

merely compute sums:

$$\phi(m) = \sum_{i=0}^m P_\lambda(i)$$

The table is limited by assuming a value of m', such that;

$$\phi(m') \doteq 1$$

The value of m' depends upon λ, but is usually small, e.g. m' = 40 for λ ≤ 20. The CBS algorithm is demonstrated using Fortran routines, IPTABL and POISON.

Subroutine IPTABL figure 7 returns the one probability and 2<sup>L</sup>-2 conditional probabilities in the proper order with P(X<sub>1</sub>=1) stored at dummy location TABLE(1). An example coding of IPTABL is given for the Poisson distribution where λ is passed in FL. The total number of bits to be generated by table compares is passed in dummy argument L. ERR=0 indicates normal termination of IPTABL while ERR=-1 signals an error. The value in TABLE where the error occurred is also replaced by -1.

Probabilities between arbitrary limits are computed with function subprogram POISON(Poisson). Minus one is returned for unacceptable arguments figure 8.

#### SUMMARY

Two special purpose instructions microprogrammed to perform the GFSR and CBS algorithms result in a wide variety of sampling possibilities merely by providing tables. The tables are generated once for each distribution required and remain unchanged (read-only). Greater speed is obtained, however, if the tables are set in control memory, since for a L bit CBS sample, L + 1 memory fetches are required from the conditional probability table and 2L + 2 fetches and L + 1 stores are required in the GFSR routine. Altogether 4(L + 1)

```

SUBROUTINE IPTABL(L, TABLE, ERR, FL)
C L=TOTAL NUMBER OF BITS SAMPLED USING TABLE.
C J2=2**J IN ITERATIVE FORM, USED TO INDEX TABLE.
C MM=2**(L-J) IN ITERATIVE FORM.
C JC=BIT COUNTER, COUNTS TO L.
  DIMENSION TABLE(1)
  ERR=0
  MM=2**(L+1)
  J2=1
  DO 1 JJ=1,L
    JC=JJ-1
    MM=MM/2
C II=PERMUTATION COUNTER.
  DO 2 II=1,J2
C OBTAIN BIT COMBINATION IN INTEGER FORM.
    I=II-1
C NORMALIZE BIT COMBINATION TO GET IXBAR(J).
    IXBAR=I*MM
C CALCULATE UPPER AND LOWER LIMITS.
    IHI=IXBAR+MM-1
    ILOW=IXBAR+MM/2
    TOP=POISON(ILOW, IHI, FL)
    BOT=POISON(IXBAR, IHI, FL)
    IF(TOP.LT.O.OR.BOT.LT.O) GO TO 4
C CALCULATE TABLE ENTRY.
    IF(BOT.GT.O) GO TO 3
    TABLE(I+J2)=0
    GO TO 2
4    TABLE(I+J2)=-1.
    ERR=-1.
    GO TO 2
3    TABLE(I+J2)=TOP/BOT
2    CONTINUE
    J2=2*J2
1    CONTINUE
    RETURN
  END

```

Fig. 7--FORTRAN subroutine for computing conditional probability table of CBS algorithm and POISSON distribution.

```

FUNCTION POISON(ILOWER, IUPPER, FL)
C POISON = SUM ON I FROM ILOWER TO IUPPER OF
C EXP(-FL)*FL**I/I FACTORIAL.
  IF(FL.LE.O.OR.ILOWER.LT.O.OR.
  XIUPPER.LT.O) GO TO 3
  IF(ILOWER.GT.IUPPER) GO TO 1
  IUP=IUPPER+1
  ILOW=ILOWER+1
  PX=EXP(-FL)
  SUM=0
  DO 2 I=1, IUP
    IF(I.EQ.ILOW) SUB=SUM
    SUM=SUM+PX
2    PX=PX*FL/I
    POISON=SUM-SUB
  RETURN
1    POISON=0
  RETURN
3    POISON=-1.
  RETURN
END

```

Fig. 8--FORTRAN function for computing POISSON cumulative distribution function,  $\phi(m)$ .



memory references are required compared to half that many when control memory is used for TAB2. Whether or not control memory is used for table storage depends upon the machine's architecture and memory space.

Higher level simulation languages are appropriate for use with these microprogrammed instructions. For example, a simulation language (SL ) compiler easily generates the corresponding conditional probability table and issues a VAR instruction. The compiler may even determine table length,  $L$ , by observing convergence to  $P(X_{j+1}|X_j) \approx 0.5$ . Greater resolution may be obtained by augmentation with a uniform random number from RND.

Dynamic vector allocation would allow creation and deletion of probability tables as needed. Storage is conserved but the tables may have to be generated again. Details and programs for generating CBS tables are given in reference 8.

The CBS algorithm is completely general, and because of the speed of GFSR and control memory, highly desirable. Software sampling routines may require far fewer random numbers, but this becomes less important with rapid means of producing random numbers. Microprogramming of present software routines is impractical because of their diversity. Thus, a general algorithm which fits a wide class of sampling problems is ideal for microprogramming.

Finally, little has been said about the arbitrary sense of CBS, but it should be noted that any distribution that can be generated by a digital computer can be generated by the microprogrammed CBS algorithm. Therefore, empirical or experimental distributions, represented by histograms, may also be applied merely by supplying a conditional probability table.

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